The effect of road network patterns on pedestrian safety:

A Zone-based Bayesian spatial modeling approach

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ABSTRACT

Pedestrian safety is increasingly recognized as a major public health concern. Extensive safety studies have been conducted to examine the influence of multiple variables on the occurrence of pedestrian–vehicle crashes. However, the explicit relationship between pedestrian safety and road network characteristics remains unknown. This studies particularly focused on the role of different road network patterns on the occurrence of crashes involving pedestrians. A global integration index via space syntax was introduced to quantify the topological structures of road networks. The Bayesian Poisson-lognormal (PLN) models with conditional autoregressive (CAR) prior were then developed via three different proximity structures: contiguity, geometry-centroid distance, and road network connectivity. The models were also compared with the PLN counterpart without spatial correlation effects. The analysis was based on a comprehensive crash dataset from 131 selected traffic analysis zones in Hong Kong.

The results indicated that higher global integration was associated with more pedestrian-vehicle crashes; the irregular pattern network was proved to be safest in terms of pedestrian crashes occurrences, whereas the grid pattern was the least safe; the CAR model with a neighborhood structure based on road network connectivity was found to outperform in model goodness-of-fit, implying the importance of accurately accounting for spatial correlation when modeling spatially aggregated crash data.

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- 36 Key words: pedestrian safety, road network patterns, global integration, road network connectivity,
- 37 zone-based approach, Bayesian spatial model

1. Introduction

As a means of primary transportation, walking holds promise for not only supporting public health, but also reducing resource and land consumption, air and noise pollution, and greenhouse gas emissions (Gomez-Ibanez et al., 2009; Ewing et al., 2011). However, walking is also associated with negative public health outcomes related with pedestrian—vehicle crashes. Approximately 400,000 pedestrians are killed in pedestrian—vehicle crashes each year worldwide (Naci et al., 2009), whereas in Hong Kong, pedestrians accounted for more than half of the road traffic fatalities (HKTD, 2014), a proportion that was much higher than that in other high-income countries (e.g., the United States, the United Kingdom, and Japan). Hence, effective road safety strategies must be formulated and implemented to improve the safety of these vulnerable road users (Wong et al., 2002; Loo et al., 2005, 2007).

Over the past decade, numerous studies have been conducted to identify the effects of potential contributing factors, such as pedestrian and driver characteristics (e.g., age and gender), road geometric features (e.g., road width, horizontal and vertical alignment), traffic characteristics (e.g., traffic volume and travel speed), neighborhood environmental conditions (e.g., urban form, land use, and access to public transport) and other factors (e.g., weather, temporal, and lighting), on the occurrence of crashes involving pedestrians (Gårder, 2004; Lee and Abdel-Aty, 2005; Kim et al., 2008; Wier et al., 2009; Ewing and Dumbaugh, 2009; Moudon et al., 2011; Siddiqui et al., 2012, 2014; Lam et al., 2014; Yao et al., 2015; Wang et al., 2016).

In addition to the preceding contributing factors, road network structures have been identified as important in determining traffic safety, especially for pedestrians (Marshall and Garrick, 2010b; Rifaat et al., 2011, 2012; Wei and Lovegrove, 2012; Dumbaugh and Zhang, 2013; Zhang et al., 2015). Urban road networks not only influence quality of life, but also promote economic development and regional growth (Florida, 2002) and improve people's physical fitness and health (Frank et al., 2003). According to Appleyard (1980), an ideal road network should possess several characteristics (such as located in a livable environment, located in neighborly territory, and being green and pleasant) in addition to acting as thoroughfares for carrying people and goods from one place to another. A reduction in pedestrian crashes is one of the road and traffic-related issues that has an effect on neighborhood life. The structure of road networks influences pedestrian safety through its effects on the fundamental characteristics of traffic volumes and journey patterns for both vehicles and pedestrians (Zhang et al., 2015). Road network patterns also affect travel activities, vehicle speeds, and driver behavior, thus influencing road safety (Quddus, 2008).

Previous studies have considered road network structures in different ways. One common way of classifying road network patterns is based on the concept of macro- and micro-level road networks proposed by Marshall (2004). Macro-level or citywide road networks distinguish roads that are typically continuous over a substantial part of the city and enable travel from one part of the city to another. In contrast, micro-level or neighborhood road networks mainly serve residential neighborhood travel, as these roads are on routes that are not continuous over a significant part of the city. Following Marshall and Garrick (2010a, 2010b, 2011), four types of citywide road network (linear, tributary, radial, and grid) and two types of neighborhood road network (tree and grid) are defined and combined to describe hierarchical road network patterns. An alternative method involves concentrating directly on the overall road network patterns of a community rather than the different types of road network pattern and then combining the different types to compose a new type. For

instance, Lovegrove and Sun (2010, 2013) classified road network patterns into five categories: fused grid, commonly used grid, cul-de-sac, three-way offset, and Dutch sustainable road safety (SRS). They found that fused grid and three-way offset patterns were safer than grid, cul-de-sac, and Dutch SRS patterns. Southworth and Ben-Joseph (2003) classified road network patterns into five other categories: gridiron, fragmented parallel, warped parallel, loops and lollipops, and lollipops on a stick. Rifaat et al. (2010, 2011, 2012) further merged fragmented parallel patterns and lollipops into a mixed pattern and concluded that relative to limited access patterns, the grid pattern was associated with more crashes and lower pedestrian crash severity.

The main limitation of the preceding studies is that their road network classification methods were largely based on the visual inspections and hence may be prone to subjectivity and uncertainty due to the unavailability of quantitative road network indices. Moreover, these manual classification approaches are time-consuming, limiting their generalization to large study regions. To overcome these problems, some researchers have attempted to extract the structures of road networks quantitatively based on topological measurements (Wang et al., 2012, 2013; Zhang et al., 2012, 2015).

This study introduced an alternative innovative road network modeling technique known as space syntax (Hillier and Hanson, 1984; Hiller, 1996) to quantify the characteristics of different road network patterns. The notion of syntax refers to relationships between different spaces, or interactions between spaces and society. It recognizes that spatial patterns or structures greatly influence human activities and behavior in urban environments. Typical applications of space syntax include pedestrian and bicyclist modeling, criminal mapping, pollution analysis, and retail viability modeling (Jiang, 1999; Ratti, 2004). Space syntax can objectively measure the effects of road network patterns on accessibility (Raford et al., 2007). This study therefore used space syntax to quantify the structures of road network patterns.

To determine the effects of road network patterns on pedestrian safety, aggregate relationships between the aforementioned independent variables and pedestrian crash counts can be established using macro-level or area-wide safety performance functions (SPFs). Traditional Poisson and Poisson-lognormal (PLN) models ignore the issue of spatial correlation, which may lead to biased parameters and incorrect inferences (Huang and Abdel-Aty, 2010). To address the spatial dependence between adjacent zones, Bayesian spatial models with a conditional autoregressive (CAR) prior have been widely applied in current road safety analysis (Quddus, 2008; Mitra, 2009; Aguero-Valverde and Jovanis, 2008, 2010; Siddiqui et al., 2012; Wang et al., 2012, 2013; Xu et al., 2014; Xu and Huang, 2015).

Within the framework of CAR, the spatial relationship is typically linked by virtue of a spatial weight matrix. One commonly used weighting scheme includes adjacency-based first-order neighbors (Miaou et al., 2003; Aguero-Valverde and Jovanis, 2008; Quddus, 2008), which can be defined as all of the observation units that connect directly with the one in question. Additional consideration is given to distance decay-based forms (Aguero-Valverde and Jovanis, 2010; Dong et al., 2014; Wang et al., 2016). However, the pure geographical proximity may be inadequate to express the spatial relationships under investigation. In reality, a more rational spatial weighting configuration is required to simultaneously consider the geographical dimension and other aspects (e.g., travel frequency, travel modes, and travel demand) that vary with the context of study. For this purpose, this study proposed a new neighborhood structure using road network connectivity to reflect the potential traffic interactions between adjacent zones.

Overall, this study described an innovative street pattern modeling technique known as space syntax to objectively quantify the structure of road networks. More specifically, the primary objective of this study was to explore how road network characteristics—in terms of global integration—are associated with zone-level pedestrian safety using both non-spatial (e.g., PLN model) and spatial (e.g., CAR model) SPFs. Another objective was to compare different spatial proximity structures at the traffic analysis zone (TAZ) level. Bayesian spatial models with a CAR prior were used with three different types of neighborhood structure: contiguity, geometry-centroid distance, and road network connectivity.

The remainder of this paper is structured as follows. Section 2 describes the data collection process. Section 3 presents the structures of non-spatial and spatial SPFs. Section 4 summarizes the results and discusses the investigation. Section 5 presents the conclusions and suggestions for the future direction of research.

2. Data

A comprehensive database relating to pedestrian—vehicle crashes and traffic-related variables for Hong Kong was established using geographical information system (GIS) techniques.

The spatial units for analysis were the TAZs from the Hong Kong Planning Vision and Strategy (PVS) zoning system, which is managed by the Hong Kong Planning Department. The PVS zoning system divides the whole of Hong Kong into different sectors and is revised every few years to better describe the zoning boundary. The PVS 338 zoning system uses a 2001-based Territorial Population and Employment Data Matrix (TPEDM). There are 338 TAZs in this zoning system. Of these, 131 TAZs (38.76%) were finally selected as having adequate traffic and geometric information. To examine the role of road network patterns and zone-related variables that contribute to pedestrian safety at a more disaggregated level, the database was organized into 4-hour periods: 07:00–11:00 (morning), 11:00–15:00 (noon), 15:00–19:00 (afternoon), 19:00–23:00 (evening), 23:00–03:00 (middle of the night), and 03:00–07:00 (dawn). Our database thus comprised 786 observations.

2.1 Measurements of road networks

Various morphological parameters of 131 TAZs in Hong Kong were derived by space syntax, including connectivity, depth, and integration. Connectivity refers to the number of lines that directly intersect a given axial line. Depth is defined as the distance in number of lines from a number of steps to the given axial line. Connectivity focuses on immediate neighborhoods, and depth focuses on k neighborhoods. Integration refers to the accessibility of a whole road network, which is derived from the mean depth.

Based on the axial map, which consists of the smallest number of longest and intersected axial lines needed to cover the whole free space and to represent the road network patterns, several variables are defined and estimated accordingly.

$$D_{i} = \sum_{d=1}^{s} d \times N_{d} = \begin{cases} connectivity & s = 1\\ local \ depth & s = k,\\ global \ depth & s = l \end{cases}$$
 (1)

where d, an integer, represents the shortest distance far from ith node, N_d denotes the number of

nodes with the shortest distance d, k is the number of neighborhoods, and l denotes the maximum shortest distance. Furthermore, it should be noted that 1 < k < l, where k is usually equal to 3 for the calculation of local depth, and the connectivity is equivalent to local depth if k = 1.

To introduce the principles of this algorithm, Fig. 1a represents a simplified road network structure. Its axial map is derived by drawing line 1, then 2, then 3 ... and the final axial map consists of 13 axial lines as shown in Fig. 1b.

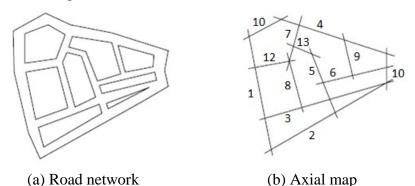


Fig. 1. A simplified road network and its axial map

Consider axial line 1 in Fig. 1b. The line directly intersects four lines, so the connectivity value of this axial line is equal to 4. The immediate neighborhoods of line 1 are lines 2, 3, 10, and 12, and their respective connectivity values are 4, 5, 2, and 3. Axial line 1 has four neighborhoods one step away, five neighborhoods two steps away, and three neighborhoods three steps away. The global depth of this axial line is thus equal to 1*4+2*5+3*3 if s=3, and the local depth with two steps away is equal to 1*4+2*5 if s=2.

 MD_i , the mean depth, is defined as the average distance from the *ith* node to other nodes. The formula for calculating mean depth is defined as

$$MD_{i} = D_{i} / (n-1) = \sum_{d=1}^{s} d \times N_{d} / (n-1) = \begin{cases} local \ mean \ depth & s=k \\ global \ mean \ depth & s=l \end{cases}.$$
 (2)

Integration measures how easily accessible the *ith* node is from other nodes in the system. The integration value is calculated by

$$I_{i} = \frac{D_{n}}{2(\text{MD}_{i} - 1)/(n - 2)} , \qquad (3)$$

where n is the number of nodes within the system and

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$$D_n = 2 \left\{ n \left\lceil \log_2 \left((n+2)/3 \right) - 1 \right\rceil + 1 \right\} / \left(n^2 - 3n + 2 \right). \tag{4}$$

Integration is a reciprocal function of MD_i , i.e., the integration value is high when the mean depth is low and vice versa. Depending on the mean depth used, i.e., local versus global, the corresponding integration value is called local versus global integration.

Continuing with this example, the global mean depth of axial line 1 is 1.9167 and the D-value is 0.2755, so the global integration of this axial line is equal to 1.6529.

In this study, connectivity, local integration, and global integration were calculated for each of the inspected 131 TAZs. Both connectivity and local integration are local parameters, as they consider immediate neighboring lines or lines with few steps. In comparison, global integration is a global parameter and considers both immediate and non-immediate neighborhoods up to a maximum number of steps away. To determine whether global integration is capable of quantitatively distinguishing different road network patterns, the network patterns of all 131 TAZs were first classified as grid, deformed grid, and irregular (Liu, 2007). These patterns are shown in Fig. 2. The values of global integration for each of these three network patterns were then calculated. The corresponding distribution is shown in Fig. 3. As illustrated, the global integration changed systematically with different road network patterns and was an appropriate measurement for describing any type of road network pattern.

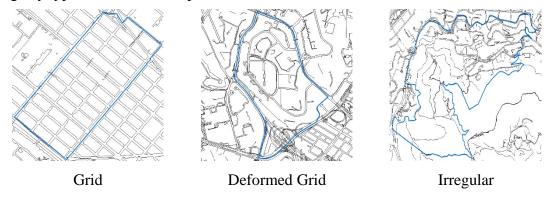


Fig. 2. Typical network patterns of Hong Kong

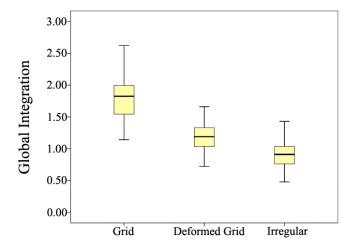


Fig. 3. Comparison of the global integration of different network patterns

2.2 Crash data

The pedestrian—vehicle crash data used in this study were obtained from the traffic information system (TIS) maintained by the Hong Kong Police Force and Transport Department. The TIS consists of three types of statistical information: crash environment profile, vehicle involvement profile, and casualty injury profile (Wong et al., 2007; Huang et al., 2016). In the TIS, all crashes are categorized according to crash severity level: fatal, serious, and slight. A fatal crash refers to a crash in which at least one person is killed or is injured and subsequently dies within 30 days of the crash. A serious crash refers to a crash in which one or more persons are injured and detained in hospital for more than 12 hours, and a slight crash is one in which one or more persons are injured, but not to the extent that a hospital stay of more than 12 hours is required. In this study, fatal and serious crashes

were classified together as killed and seriously injured (KSI) crashes (Pei et al., 2011). Fig. 4 shows the spatial distribution of aggregated crashes that occurred in the 131 TAZs of Hong Kong during

2011. Of the 2,168 pedestrian–vehicle crashes, 462 (21.31%) were KSI crashes.

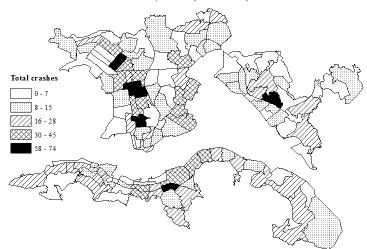


Fig. 4. Spatial distribution of pedestrian crashes in the 131 TAZs of Hong Kong

2.3 Speed data

Travel speed is believed to be a determinant factor of pedestrian crash occurrence. In this study, speed data from global positioning system (GPS) probe vehicles were used. GPS probes are present in 480 taxis traveling on the roadway network in Hong Kong. Instantaneous information on the location, speed, and travel direction of each GPS-equipped taxi is transmitted to and stored in a control center at 30-second intervals. Pei et al. (2009) adopted the journey time survey data maintained by the Department of Transportation to explore the validity of the GPS probe data and found that taxi speed was a reasonable proxy for actual travel speed. The average travel speed for each TAZ of interest was calculated for every 4-hour period using these raw data.

2.4 Exposure

Two exposure indicators—vehicle hours (VH) and pedestrian activities (PA)—were considered in the proposed models. As suggested by Pei et al. (2012), the time exposure measured by VH is more reasonable as an exposure index. To estimate the variable, the average annual daily traffic (AADT) data and its temporal (in relation to month, day of the week, and hour of the day) and directional multiplicative factors were obtained from the Hong Kong annual traffic census (ATC) system, which is maintained by the Hong Kong Transport Department (Lam et al., 2003; Tong et al., 2003). The ATC system provides the hourly traffic volume for approximately 100 core stations in Hong Kong. According to Wong and Wong's (2015, 2016) studies, a linear data projection method was adopted to combine the AADT and taxi GPS dataset. First, the scaling factors for each ATC core station were calculated. Second, the scaling factors for each TAZ were predicted according to the spatial correlation of the related core stations. Consequently, the hourly traffic spots within each TAZ could thus be estimated, and time exposure could further be estimated in relation to the last time for each spot, which was supposed to be equal to the update time of the taxi GPS data.

To calculate another exposure indicator, pedestrian trip data were first extracted from the trip-based data file (named C501A) of the Travel Characteristics Survey 2002 (TCS 2002). This

survey is conducted by the Hong Kong Transport Department about once a decade to understand the normal weekday (excluding public holidays) trip information and other information, such as the household and personal details of Hong Kong residents. The C501A file contains all trip records reported by the respondents. Two sub-files were extracted from C501A for separate analysis: one containing the data for walk-only trips, and the other containing the data for trips involving any of the 10 major mechanized transport modes. In this study, analysis of pedestrian trips data was based on walk-only trips. The PA of each inspected TAZ consisted of three parts: trip generations, trip attractions, and internal pedestrian activities.

To explore the effect of exposure on safety performance, logarithmically transformed VH and PA were incorporated into the models as explanatory factors.

2.5 Other explanatory factors

For a better estimation of the effects on pedestrian safety, the confounding factors of roadway geometric characteristics, demographics, land use patterns, and temporal distribution were controlled in the developed models.

Roadway geometric characteristics are commonly considered as contributing factors to the occurrence of crashes involving pedestrians at the TAZ level. Various macro-level geometric attributes (Noland and Quddus, 2004; Wier et al., 2009; Siddiqui et al., 2012), including the number of intersections, total road length, road density, and intersection density, are considered as possible descriptive factors. Road density is equal to road length divided by zone area per TAZ. Intersection density is defined as the number of intersections divided by road length per TAZ. In this study, road density and intersection density were included in the models.

Demographic factors were extracted from TPEDM and included different kinds of estimated values (such as the resident population, the age cohort of residents with regular residence, and the number of domestic households) in the base year listed for each TAZ. In this study, the percentage of the population aged 65 years and older was considered in the models.

Land use patterns were extracted from the outline zoning plan (OZP) maintained by the Hong Kong Town Planning Board. The OZP uses colored hatching on the base map to represent and identify different permitted land uses, such as government/institution/community, commercial, and residential. In this study, the land use pattern of each selected TAZ was estimated by the percentage of trips in each area according to the trip pattern and purpose. The land use patterns were divided into four different categories—residential, commercial, mixed (combined commercial and residential use), and others—by comparing the dominant trip purpose (more than 50%) and incorporated into the models as categorical variables, with residential as the control.

The effect of temporal distribution on pedestrian–vehicle crash occurrences by time of day was also controlled to neutralize the confounding effect of unidentified factors. For example, the six periods were included in the regression models as categorical variables with the dawn period (03:00–07:00) as the control.

2.6 Summary statistics of variables

Table 1 shows the descriptive statistics for the variables included in the final models. To ensure that all of the contributory factors included were statistically independent, the multicollinearity between them was calculated. One of the quantified factors, the variance inflation factor (VIF), was used to identify the severity of the multicollinearity in the regression model. The VIF values of all of

the independent variables were less than 5, showing no evidence of a potentially serious collinearity problem or a correlation problem affecting the regression.

Table 1 Summary Statistics of Crash Occurrences and Contributory Variables

	Min	Max	Mean	S.D.	Count	Proportion
Number of observations = 786	-	•	•	•	•	•
Total crashes	0	30	2.758	3.704		
KSI crashes	0	6	0.588	0.995		
Slight crashes	0	26	2.171	3.078		
Average speed (km/h)	7.758	77.972	26.820	12.622		
Global integration	0.479	3.624	1.146	0.428		
Road density (km/km ²)	21.198	109.392	50.074	14.962		
Intersection density (Number of intersections per km)	1.210	11.240	3.559	1.474		
Senior percentage (older than 65)	0.000	0.232	0.131	0.040		
Ln(PA)	0.000	6.238	3.153	1.819		
Ln(VH)	5.371	13.516	11.044	1.250		
07:00-11:00					131	16.7%
11:00-15:00					131	16.7%
15:00–19:00					131	16.7%
19:00-23:00					131	16.7%
23:00-03:00					131	16.7%
03:00-07:00					131	16.7%
Residential					462	58.7%
Commercial					150	19.1%
Mixed					102	13.0%
Others					72	9.2%

3. Methodology

For the purpose of comparison, two types of SPFs—non-spatial and spatial models—were established and estimated using the Markov chain Monte Carlo (MCMC) approach and full Bayesian method (Gelman et al., 2014), respectively. Both of these modeling techniques are briefly presented as follows.

3.1 Non-spatial models

Crash data are nonlinear, discrete, and non-negative data that approximately follow the Poisson distribution (Lord and Mannering, 2010). One underlying assumption of the Poisson model is that the mean of the data is restricted to be equal to the variance, i.e., $E(y_i) = Var(y_i)$. If this equality does not hold, the data are said to be under-dispersed $(E(y_i) > Var(y_i))$ or over-dispersed $(E(y_i) < Var(y_i))$, and the coefficient's vector is biased if that is not taken into account. In most practical cases, the variance of pedestrian crash data is greater than the mean, so the data are considered to be over-dispersed (Noland and Quddus, 2004). Over-dispersion can be resolved by

assuming that the crash data are distributed in a negative binomial (NB) or PLN distribution. To test the appropriateness of using the traditional count models (e.g., NB) rather than zero-inflated models (e.g., zero-inflated negative binomial (ZINB)), the Vuong test was conducted to compare the predictive probability of two candidate models (Vuong, 1989). Finally, the Voung statistic for testing the ZINB versus the normal-count NB was computed as -3.421, which was less than -1.96 and favored the traditional NB model. Although the NB model has been widely applied in road safety analysis (Washington et al., 2011), Lord and Miranda-Moreno (2008) recently found that the PLN model could be a better alternative in cases of small sample size and low sample mean values.

Consequently, the PLN distribution is preferred over other mixed-Poisson distributions by including the lognormal error distribution in the Poisson distribution. The PLN model (Model 1) is specified as follows:

$$y_i \mid \lambda_i \sim \text{Poisson}(\lambda_i), \tag{5}$$

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$$\lambda_i = \exp\left(\beta_0 + \beta_1 \ln\left(VH_i\right) + \beta_2 \ln(PA_i) + \sum_{k=3}^p \beta_k X_{ik} + \theta_i\right), \tag{6}$$

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$$\theta_i \sim \text{Normal}(0, \tau_\theta^2),$$
 (7)

where y_i is the number of crashes for TAZ_i per period, λ_i is the expected number of crashes per period, β_0 is the intercept term, $\beta = (\beta_1, \beta_2, \dots, \beta_p)'$ is a regression coefficient vector, VH_i and PA_i are the exposure variables of TAZ_i , X_i is a vector of explanatory variables, θ_i is the unobserved (i.e., uncorrelated) heterogeneity error component of the model, and τ_{θ}^2 is the precision (inverse of the variance) parameter with a prior gamma distribution Ga(0.5, 0.0005), as suggested by Wakefield et al. (2000) and used by Aguero-Valverde and Jovanis (2008, 2010).

21 3.2 Spatial models

The PLN model is able to capture unstructured over-dispersion or unobserved heterogeneity caused by omitted variables. However, this model ignores the issue of spatial dependence among spatial TAZs, which violates the traditional Gauss-Markov assumption involved in regression modeling (Huang and Abdel-Aty, 2010). To account for the heterogeneity that occurs due to spatial correlation, a second random effect component ϕ_i is added to the model, and Equation 6 can be rewritten as

$$\lambda_{i} = \exp\left(\beta_{0} + \beta_{1} \ln\left(VH_{i}\right) + \beta_{2} \ln(PA_{i}) + \sum_{k=3}^{p} \beta_{k} X_{ik} + \theta_{i} + \phi_{i}\right), \tag{8}$$

29 where ϕ_i is the spatial correlation error component.

- The effect of spatial correlation, ϕ_i , is assigned as intrinsic CAR with a distribution $N(\overline{\phi_i}, \tau_i^2)$
- 2 as recommended by Besag et al. (1991), where

$$\overline{\phi}_{i} = \frac{\sum_{i \neq j} \omega_{ij} \phi_{j}}{\sum_{i \neq j} \omega_{ij}} \quad \text{and} \quad \tau_{i}^{2} = \frac{\tau_{\phi}^{2}}{\sum_{i \neq j} \omega_{ij}}, \tag{9}$$

- 4 in which ω_{ij} is the adjacency index for TAZ_i and TAZ_j in the spatial weight matrix W and
- 5 au_ϕ^2 is also assumed a gamma prior with $ext{ Ga} ig(0.5, 0.0005 ig)$ as suggested by Wakefield et al. (2000).
- To determine the contribution of spatial correlation, the proportion of variability in the random effect components is calculated as

$$\alpha = \frac{\mathrm{SD}(\phi)}{\mathrm{SD}(\phi) + \mathrm{SD}(\theta)}.\tag{10}$$

- 9 where SD is the empirical marginal standard deviation function.
- As there is no prior information available, a non-information prior distribution for β is
- specified as Normal(0, 100,000). A uniform prior distribution is assumed for β_0 .
- 12 3.3 Proximity matrix
- In the CAR models, the specification of spatial weight matrix W is a major concern because
- the choice of proximity matrix may have a substantive effect on the results. In the present study,
- 15 three different forms of neighborhood structure constructed by contiguity (Model 2a),
- 16 geometry-centroid distance (Model 2b), and road network connectivity (Model 2c) were used.
- 17 Model 2a proximity structure based on contiguity
- The simplest form for W is a "0/1" contiguity matrix, wherein $\omega_{ij} = 1$ if TAZ_i and TAZ_j
- are spatially adjacent and 0 otherwise.
- 20 Model 2b proximity structure based on geometry-centroid distance
- A more general setup uses distance to reflect the correlation of TAZs. An alternative form of W,
- defining ω_{ii} as a function of the distance between TAZs, is

$$\omega_{ij} = C\left(d_{ij}\right),\tag{11}$$

- 24 where d_{ij} is the geometry-centroid distance between TAZ_i and TAZ_j , and $C(d_{ij})$ is usually a
- decreasing function of d_{ij} so that the TAZs close to each other have a large covariance. In this study,
- 26 the exponential function $C(d_{ij}) = \exp(-d_{ij}/\omega)$, where ω is a spatial decay parameter, was used to

- calculate the spatial weight matrix. According to a previous study by Carlin and Ma (2007), ω was set to the average distance among TAZs in this study.
- 3 Model 2c proximity structure based on road network connectivity
- In this study, another alternative approach to defining the spatial weights matrix W was based on road network connectivity. The general expression for W is provided by

$$\omega_{ii} = Q(e_{ii}) * C(d_{ii}), \tag{12}$$

- 7 where e_{ij} represents the links connecting TAZ_i and TAZ_j . The links are defined as the road
- 8 segments between two TAZs. $Q(e_{ij})$ is the traffic volume of the links. The total number of road
- 9 segments between two TAZs is calculated as a quantitative index to represent the traffic volume of
- 10 links. $C(d_{ij})$ is a negative exponential function of the geometry-centroid distance d_{ij} , i.e.,
- 11 $C(d_{ij}) = \exp(-d_{ij}/\omega)$. ω was similarly set to the average distance among TAZs.
- 12 3.4 Model comparison
- To compare the different models, the deviance information criterion (DIC), an extension of the
- 14 Akaike information criterion (AIC) proposed by Spiegelhalter et al. (2002), was used. The DIC
- provides a Bayesian measure of model complexity and fitting. Based on the deviance statistic $D(\theta)$,
- the criterion DIC is defined as

DIC=
$$\overline{D(\theta)} + p_D$$
, (13)

- where $\overline{D(\theta)}$ is the posterior expectation deviance that can be a Bayesian measure of fitting, i.e.,
- 19 $\overline{D(\theta)} = E(D(\theta)|y)$. p_D is the effective number of parameters in relation to a model and is defined
- 20 as

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$$p_{D} = E(D(\theta)|y) - D(E(D(\theta)|y)). \tag{14}$$

22 Similar to the AIC, models with lower DIC values are preferred.

4. Results and discussion

Depending on the complexity of the models, three MCMC chains from different starting points were estimated to ensure convergence. The first 5,000 iterations in each chain were discarded as burn-ins. A further 5,000 iterations were then performed for each chain, resulting in a sample distribution of 15,000 for each parameter. Convergence of the models was monitored by visual examination of the MCMC chains, autocorrelation plots, and Gelman-Rubin statistic plots.

4.1 Model comparison

The model performance results are shown in Table 2. The model performance results are shown in Table 2. The posterior means for the standard deviations of spatial correlation ϕ for all CAR models are statistically significant, suggesting that the observed pedestrian-vehicle crashes are spatially correlated among the TAZs. Similar to ϕ , the posterior means for the standard deviations of uncorrelated heterogeneity θ are also statistically significant.

As shown by the DIC values, the CAR models substantially outperform their PLN counterpart. This result is not surprising to some extent, as the CAR models not only account for the unstructured over-dispersion, but also take the spatial correlation between adjacent TAZs into account.

By comparing the DIC values of CAR models (DIC = 2,576.49 for Model 2a, 2,564.74 for Model 2b, and 2,559.17 for Model 2c), it can be seen that the consideration of proximity structure based on road network connectivity among adjacent TAZs is more reasonable. This result is expected, as both Model 2a and Model 2b only consider contiguity or distance relationships. In contrast, Model 2c is a quantitative measurement of transport infrastructures, which actually concentrates on transport activities (e.g., vehicular trips and pedestrian trips). For example, the relationships between entities that are geographically connected to neighbors are likely to be loose, as neighboring entities have low road network connectivity, resulting in a total number of vehicular and pedestrian trip productions and attractions that is small or very close to zero.

In summary, it is essential to identify and explore the spatial correlation of pedestrian—vehicle crashes. Moreover, the CAR model with a neighborhood structure based on road network connectivity can properly account for the spatial heterogeneity between the observations.

Table 2 Model Comparison

Model	PLN model	CAR models					
Model	Model 1	Model 2a	Model 2b	Model 2c			
S.D. of θ	0.364 (0.297, 0.434)*	0.132 (0.027, 0.260)	0.127 (0.028, 0.253)	0.114 (0.025, 0.228)			
S.D. of ϕ	-	0.415 (0.344, 0.495)	0.389 (0.325, 0.456)	0.394 (0.328, 0.463)			
α	-	0.769 (0.602, 0.938)	0.764 (0.598, 0.935)	0.783 (0.625, 0.941)			
DIC	2,609.870	2,576.490	2,564.740	2,559.170			

^{*} Estimated mean (95% Bayesian credible interval) for the parameter

4.2 Parameter estimation

The estimation results of Model 2c are presented in Table 3. The estimation results of Model 2c are presented in Table 3. The main objective of this study was to explore the effect of road network structures on the occurrence of pedestrian-vehicle crashes. For this purpose, a global integration index with space syntax was derived and used. As shown in Table 3, an increase in global integration increases the likelihood of a pedestrian-vehicle crash occurring at the 5% significance level (mean = 0.255; 95% CI = [0.090, 0.424]). A road network with higher global integration undoubtedly indicates roads that provide greater accessibility to a large number of other roads acting as conduits for more origin/destination pairs. These networks consequently experience more use than do less connected and less accessible networks. A higher global integration also indicates that a road within a network is likely to include a greater number of connections (intersections) to reach other road. As

the number of intersections increases, so does the probability of complex vehicular maneuvers. In addition, a network with higher global integration tends to have more complicated routes, which increases the potential for pedestrian—vehicle conflicts. To further illustrate the relationship between road networks and global integration, the values of global integration for each TAZ were calculated. According to the three studied network patterns, nine samples from one hundred and thirty-one TAZs were chosen. Each category contained three samples. The global integration of the road network for each sample zone is shown in Table 4.

Table 3 Modeling Results of Model 2c

Variable	Posterior Mean	Posterior S.D.	(95% CI)		
Constant	-3.186*	0.364	(-3.920,-2.475)		
Global integration	0.255*	0.087	(0.090, 0.424)		
Road density	0.002	0.002	(-0.002, 0.007)		
Intersection density	0.088*	0.028	(0.034, 0.142)		
Average speed	-0.021*	0.003	(-0.028, -0.015)		
Senior percentage	0.568	1.045	(-1.474, 2.550)		
Commercial	-0.237*	0.119	(-0.470, -0.003)		
Mixed	0.574*	0.108	(0.365, 0.789)		
Others	-0.660*	0.207	(-1.076, -0.260)		
Residential	(control)				
07:00-11:00	0.959*	0.155	(0.667, 1.282)		
11:00-15:00	0.947*	0.162	(0.643, 1.284)		
15:00–19:00	1.144*	0.153	(0.849, 1.460)		
19:00-23:00	0.998*	0.142	(0.731, 1.289)		
23:00-03:00	0.400*	0.163	(0.083, 0.726)		
03:00-07:00	(control)				
Ln(PA)	0.235*	0.033	(0.168, 0.298)		
Ln(VH)	0.168*	0.035	(0.108, 0.242)		

^{*} Statistically significant at the 5% level.

Table 4 Global Integration of Nine Samples

	G1	G2	G3	D1	D2	D3	I1	I 2	13
Network				闡	AND				THE THE
Integration	3.624	1.997	1.773	0.915	1.046	0.909	0.768	0.955	0.891

Note: G = grid; $D = deformed\ grid$; I = irregular.

The results presented in Table 4 show that a greater global integration value means the greater likelihood of a grid pattern. In addition, the grid pattern network is more compact because road density tends to be higher with higher global integration. The safety of the studied network patterns

generally improved as global integration decreased, with a grid being the least safe pattern and an irregular pattern the safest. This finding is similar to that of Wang et al. (2013). Deformed grid and irregular design patterns generally have less connectivity, frequent curves, and loops. The grid pattern generally has straight roadways, which makes the process of driving comfortable and predictable. However, the favorable driving environment on the grid pattern often increases fatigue in drivers and reduces their attention, which may enhance the possibility of pedestrian—vehicle crashes occurring in critical situations (Rifaat et al., 2011, 2012). In contrast, the presence of frequent curves and loops in the deformed grid and irregular design patterns may provoke drivers to be more cautious. Further, given that the irregular pattern may restrict drivers' sight distance and make vehicle control more difficult, drivers are likely to compensate with increased attention (Hedlund, 2000; Levym and Miller, 2000; McCarthy and Talley, 1999). Moreover, the grid network pattern has higher potential for pedestrian—vehicle conflict points and less traffic calming than the deformed grid and irregular design patterns.

Consistent with previous studies by Ewing and Dumbaugh (2009) and Siddiqui et al. (2012), our results show that increased intersection density is associated with a higher chance of pedestrian–vehicle crashes at the 5% significance level (mean = 0.088; 95% CI = [0.034, 0.142]). Given that the intersection density is defined by the number of intersections divided by road length per TAZ, more intersections and shorter road lengths in particular TAZs result in higher intersection density. The higher intersection density of a given TAZ increases pedestrian exposure with a higher probability of pedestrian–vehicle interactions, as intersections are often locations with a relatively high number of pedestrians crossing streets and with unexpected driver behavior.

The effect of average speed on the occurrence of pedestrian–vehicle crashes was also explored. The results imply that average speed reduces the likelihood of a pedestrian–vehicle crash occurring at the 5% significance level (mean = -0.021; 95% CI = [-0.028, -0.015]). Contrary to the findings of Gårder (2004) and Cuadrado et al. (2008), the results show that higher average speed is associated with a lower pedestrian crash occurrence. The reduction in the number of pedestrian–vehicle crashes in zones with higher average speeds may be due to favorable built environments in which road network design features (such as separation of pedestrians from vehicles by time or space, and measures that increase the visibility and conspicuity of pedestrians) are more reasonable and road infrastructures (such as pedestrian refuge islands, roadway lighting, overpasses, and underpasses) are relatively good at reducing the exposure of pedestrians to potential harm, both on the roadside and when crossing (Garber and Gadirau, 1998).

Four land use patterns were examined to identify their influence on the occurrence of crashes involving pedestrians: residential, commercial, mixed, and others. Compared with residential areas, the likelihood of a pedestrian-vehicle crash occurring in the commercial areas (mean = -0.237; 95% CI = [-0.470, -0.003]) and other areas (mean = -0.660; 95% CI = [-1.076, -0.260]) of Hong Kong is significantly reduced at the 5% level, and the likelihood of a pedestrian-vehicle crash occurring in mixed areas (mean = 0.574; 95% CI = [0.365, 0.789]) is significantly increased at the 5% level. Contrary to the findings of previous studies by Kim et al. (2008) and Wier et al. (2009), the reduction of pedestrian-vehicle crash occurrences in commercial and other areas of Hong Kong may be due to the presence of additional overpasses. Fig. 5 illustrates two figures comparing the commercial and residential areas during the weekday afternoon off-peak period on the same day. It can be seen that vehicle flow and travel speed in commercial areas may be higher than in the residential areas throughout the day. However, overpasses in commercial areas usually provide separate pedestrian

linkages to reduce the chances of conflicts compared with the level crossings in residential areas. Moreover, in most of the residential areas of Hong Kong, ground-floor shops are popular and more shops are generally associated with more attractions, restaurants, and dwelling units. This could eventually increase the probability of pedestrian—vehicle crashes. The relationship between pedestrian crashes and mixed areas has been supported by the finding of Ewing and Dumbaugh (2009) that pedestrians in neighborhoods with mixed land use patterns have more difficulty perceiving the environment as safe than do their counterparts in neighborhoods dominated by monotonous land use patterns.





Commercial area

Residential area

Fig. 5. Comparison between commercial and residential areas

The effect of time of day on the occurrence of pedestrian-related crashes was also investigated. A day was divided into six categories: 07:00–11:00 (morning), 11:00–15:00 (noon), 15:00–19:00 (afternoon), 19:00-23:00 (evening), 23:00-03:00 (middle of the night), and 03:00-07:00 (dawn). Compared with the dawn period, the temporal distribution in the models represents the crash heterogeneity at different times of day. The afternoon off-peak period from 15:00 to 19:00 has the highest probability of pedestrian–vehicle crashes (mean = 1.144; 95% CI = [0.849, 1.460]). Although the traffic flow in this period is lower than that in the peak period, the number of daily activities (leisure, shopping, meetings, etc.) may be higher. The second and third highest probabilities of pedestrian-vehicle crashes occur in the evening peak period (19:00–23:00) (mean = 0.998; 95% CI = [0.731, 1.289]) and morning peak period (07:00-11:00) (mean = 0.959; 95% CI = [0.667, 1.282]). During the peak periods, each TAZ is expected to have higher internal and external pedestrian activities, for example, morning and evening walks, which increase the probability of pedestrian-related crashes. The period with the fourth highest probability of accidents was the noon off-peak period from 11:00-15:00 (mean = 0.947; 95% CI = [0.643, 1.284]). As more workers go to restaurants for meals during this period, there may be an increase in pedestrian and vehicle trips. In turn, this may increase pedestrian vulnerability to crashes. The period with the lowest probability of accidents happening is the midnight period (23:00-03:00) (mean = 0.400; 95% CI = [0.083, 0.726]). Although there are extremely low traffic flows during the midnight period, both drivers and pedestrians may engage in higher risk behavior in poor lighting conditions, such as driving while fatigued, aggressive driving, and driving under the influence of alcohol/drugs, which leads to more pedestrian crashes. This finding is consistent with the results of previous studies by Lee and

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8 9 Abdel-Aty (2005) and Moudon et al. (2011).

An increase in log-transformed VH significantly increases the likelihood of a pedestrian—vehicle crash occurring at the 5% level (mean = 0.168; 95% CI = [0.108, 0.242]). This finding shows that crash occurrences increase with increases of VH. The higher probability of pedestrian—vehicle interactions due to increased vehicular traffic volumes explains why log-transformed VH is positively correlated with the occurrence of crashes involving pedestrians. An increase in logarithmically transformed PA (mean = 0.235; 95% CI = [0.168, 0.298]) also significantly increases the likelihood of a pedestrian—vehicle crash occurring at the 5% level. This finding is consistent with previous studies by Lam et al. (2014) and Yao et al. (2015). The higher log-transformed PA of particular TAZs is highly correlated with more consequent pedestrian activities (compulsory trips, shopping trips, escort trips, etc.). A higher volume of pedestrian activities is expected to increase the possibility of pedestrian—vehicle crashes.

5. Conclusions

The present study attempted to investigate the role of road network patterns and zone-related variables in zone-level pedestrian safety analysis. To determine the explicit relationship between road network characteristics and pedestrian safety, three topological indices (connectivity, local integration, and global integration) of road network patterns via space syntax were derived and used to quantitatively represent the structures of road networks for each TAZ. Global integration outperformed the other two local measurements due to its ability to effectively quantify the different road network patterns.

The relationships between pedestrian–vehicle crashes and various contributory factors were developed, compared, and analyzed using both non-spatial PLN (Model 1) and spatial CAR models with different neighborhood structures (Models 2a–2c). The results showed that the CAR models outperformed the PLN model in terms of goodness-of-fit, indicating that spatial correlation should be addressed when modeling pedestrian–vehicle crashes for spatially aggregated data. Moreover, the model with the proximity structure based on road network connectivity (Model 2c) was found to have significantly better model performance than those based solely on contiguity (Model 2a) and geometry-centroid distance (Model 2b). This neighborhood structure used the total number of road segments as a quantitative index of traffic volumes to reflect the connections between TAZs. This configuration was appealing in the case of neighboring zones with different vehicular traffic and pedestrian volumes. The goal was to determine the geographical adjacency and traffic association simultaneously.

A significant positive relationship between global integration and crashes involving pedestrians could be observed, meaning higher global integration was related to more pedestrian—vehicle crashes. A greater global integration value clearly revealed closer associations between the connections within road networks and was correlated with more potential for pedestrian—vehicle conflicts. Among the three mentioned road network patterns, the grid pattern network had the highest global integration value and the irregular pattern network had the smallest value. Thus, a grid pattern network was the least safe, a deformed grid pattern network was the second safest, and an irregular pattern network was the safest.

For future study, one direction is to conduct spatial analysis for both larger-scale (e.g., broad districts of Hong Kong) and smaller (e.g., census tracts) area aggregations. In some applications there is a need to consider different scales of aggregation within analysis. Scale change issues are

often considered modifiable areal unit problems that may lead to biased inferences.

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